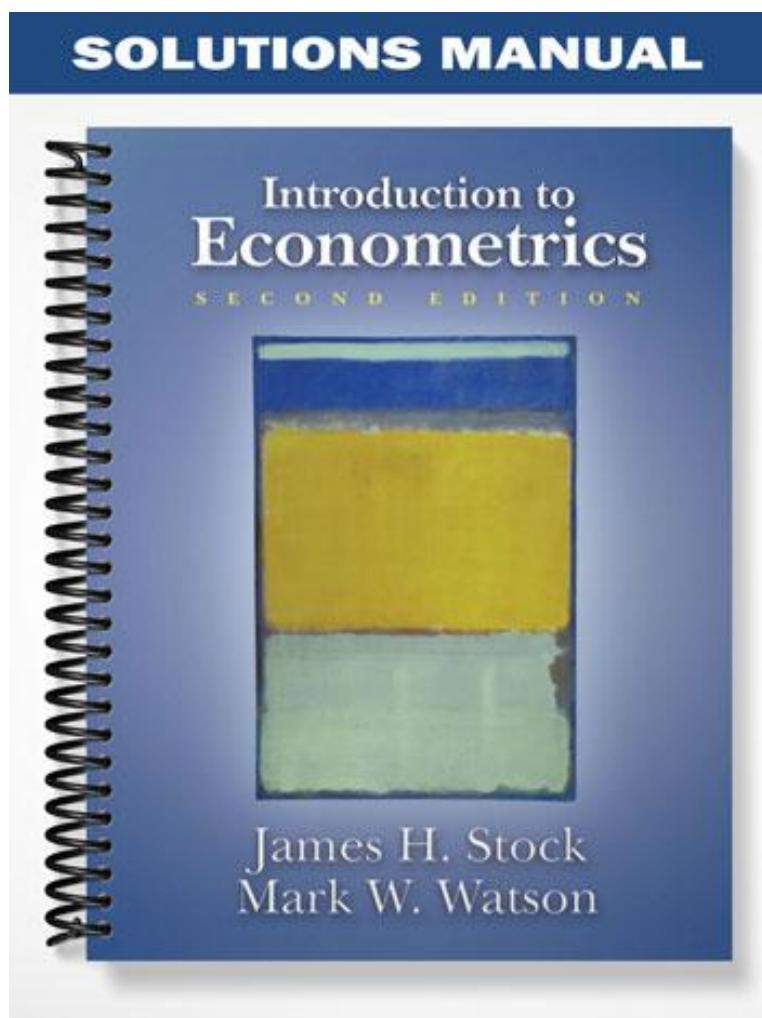


SOLUTIONS MANUAL



PART ONE

Solutions to Exercises

Chapter 2

Review of Probability

■ Solutions to Exercises

1. (a) Probability distribution function for Y

Outcome (number of heads)	$Y = 0$	$Y = 1$	$Y = 2$
probability	0.25	0.50	0.25

- (b) Cumulative probability distribution function for Y

Outcome (number of heads)	$Y < 0$	$0 \leq Y < 1$	$1 \leq Y < 2$	$Y \geq 2$
Probability	0	0.25	0.75	1.0

(c) $\mu_Y = E(Y) = (0 \times 0.25) + (1 \times 0.50) + (2 \times 0.25) = 1.00$

Using Key Concept 2.3: $\text{var}(Y) = E(Y^2) - [E(Y)]^2$, and

$$E(Y^2) = (0^2 \times 0.25) + (1^2 \times 0.50) + (2^2 \times 0.25) = 1.50$$

so that $\text{var}(Y) = E(Y^2) - [E(Y)]^2 = 1.50 - (1.00)^2 = 0.50$.

2. We know from Table 2.2 that $\Pr(Y = 0) = 0.22$, $\Pr(Y = 1) = 0.78$, $\Pr(X = 0) = 0.30$,

$\Pr(X = 1) = 0.70$. So

- (a)

$$\begin{aligned} \mu_Y &= E(Y) = 0 \times \Pr(Y = 0) + 1 \times \Pr(Y = 1) \\ &= 0 \times 0.22 + 1 \times 0.78 = 0.78, \end{aligned}$$

$$\begin{aligned} \mu_X &= E(X) = 0 \times \Pr(X = 0) + 1 \times \Pr(X = 1) \\ &= 0 \times 0.30 + 1 \times 0.70 = 0.70. \end{aligned}$$

- (b)

$$\begin{aligned} \sigma_x^2 &= E[(X - \mu_X)^2] \\ &= (0 - 0.70)^2 \times \Pr(X = 0) + (1 - 0.70)^2 \times \Pr(X = 1) \\ &= (-0.70)^2 \times 0.30 + 0.30^2 \times 0.70 = 0.21, \end{aligned}$$

$$\begin{aligned} \sigma_Y^2 &= E[(Y - \mu_Y)^2] \\ &= (0 - 0.78)^2 \times \Pr(Y = 0) + (1 - 0.78)^2 \times \Pr(Y = 1) \\ &= (-0.78)^2 \times 0.22 + 0.22^2 \times 0.78 = 0.1716. \end{aligned}$$

- (c) Table 2.2 shows $\Pr(X = 0, Y = 0) = 0.15$, $\Pr(X = 0, Y = 1) = 0.15$, $\Pr(X = 1, Y = 0) = 0.07$, $\Pr(X = 1, Y = 1) = 0.63$. So

$$\begin{aligned}\sigma_{XY} &= \text{cov}(X, Y) = E[(X - \mu_X)(Y - \mu_Y)] \\ &= (0 - 0.70)(0 - 0.78)\Pr(X = 0, Y = 0) \\ &\quad + (0 - 0.70)(1 - 0.78)\Pr(X = 0, Y = 1) \\ &\quad + (1 - 0.70)(0 - 0.78)\Pr(X = 1, Y = 0) \\ &\quad + (1 - 0.70)(1 - 0.78)\Pr(X = 1, Y = 1) \\ &= (-0.70) \times (-0.78) \times 0.15 + (-0.70) \times 0.22 \times 0.15 \\ &\quad + 0.30 \times (-0.78) \times 0.07 + 0.30 \times 0.22 \times 0.63 \\ &= 0.084,\end{aligned}$$

$$\text{cor}(X, Y) = \frac{\sigma_{XY}}{\sigma_X \sigma_Y} = \frac{0.084}{\sqrt{0.21 \times 0.1716}} = 0.4425.$$

3. For the two new random variables $W = 3 + 6X$ and $V = 20 - 7Y$, we have:

(a)

$$\begin{aligned}E(V) &= E(20 - 7Y) = 20 - 7E(Y) = 20 - 7 \times 0.78 = 14.54, \\ E(W) &= E(3 + 6X) = 3 + 6E(X) = 3 + 6 \times 0.70 = 7.2.\end{aligned}$$

(b)

$$\begin{aligned}\sigma_W^2 &= \text{var}(3 + 6X) = 6^2 \cdot \sigma_X^2 = 36 \times 0.21 = 7.56, \\ \sigma_V^2 &= \text{var}(20 - 7Y) = (-7)^2 \cdot \sigma_Y^2 = 49 \times 0.1716 = 8.4084.\end{aligned}$$

(c)

$$\begin{aligned}\sigma_{WV} &= \text{cov}(3 + 6X, 20 - 7Y) = 6(-7)\text{cov}(X, Y) = -42 \times 0.084 = -3.528 \\ \text{cor}(W, V) &= \frac{\sigma_{WV}}{\sigma_W \sigma_V} = \frac{-3.528}{\sqrt{7.56 \times 8.4084}} = -0.4425.\end{aligned}$$

4. (a) $E(X^3) = 0^3 \times (1-p) + 1^3 \times p = p$

(b) $E(X^k) = 0^k \times (1-p) + 1^k \times p = p$

(c) $E(X) = 0.3$

$$\text{var}(X) = E(X^2) - [E(X)]^2 = 0.3 - 0.09 = 0.21$$

$$\text{Thus, } \sigma = \sqrt{0.21} = 0.46.$$

To compute the skewness, use the formula from exercise 2.21:

$$\begin{aligned}E(X - \mu)^3 &= E(X^3) - 3[E(X^2)][E(X)] + 2[E(X)]^3 \\ &= 0.3 - 3 \times 0.3^2 + 2 \times 0.3^3 = 0.084\end{aligned}$$

$$\text{Alternatively, } E(X - \mu)^3 = [(1 - 0.3)^3 \times 0.3] + [(0 - 0.3)^3 \times 0.7] = 0.084$$

$$\text{Thus, skewness} = E(X - \mu)^3 / \sigma^3 = 0.084 / 0.46^3 = 0.87.$$

To compute the kurtosis, use the formula from exercise 2.21:

$$\begin{aligned} E(X - \mu)^4 &= E(X^4) - 4[E(X)][E(X^3)] + 6[E(X)]^2[E(X^2)] - 3[E(X)]^4 \\ &= 0.3 - 4 \times 0.3^2 + 6 \times 0.3^3 - 3 \times 0.3^4 = 0.0777 \end{aligned}$$

$$\text{Alternatively, } E(X - \mu)^4 = [(1 - 0.3)^4 \times 0.3] + [(0 - 0.3)^4 \times 0.7] = 0.0777$$

$$\text{Thus, kurtosis is } E(X - \mu)^4 / \sigma^4 = 0.0777 / 0.46^4 = 1.76$$

5. Let X denote temperature in °F and Y denote temperature in °C. Recall that $Y = 0$ when $X = 32$ and $Y = 100$ when $X = 212$; this implies $Y = (100/180) \times (X - 32)$ or $Y = -17.78 + (5/9) \times X$. Using Key Concept 2.3, $\mu_x = 70$ °F implies that $\mu_y = -17.78 + (5/9) \times 70 = 21.11$ °C, and $\sigma_x = 7$ °F implies $\sigma_y = (5/9) \times 7 = 3.89$ °C.
6. The table shows that $\Pr(X = 0, Y = 0) = 0.045$, $\Pr(X = 0, Y = 1) = 0.709$, $\Pr(X = 1, Y = 0) = 0.005$, $\Pr(X = 1, Y = 1) = 0.241$, $\Pr(X = 0) = 0.754$, $\Pr(X = 1) = 0.246$, $\Pr(Y = 0) = 0.050$, $\Pr(Y = 1) = 0.950$.

(a)

$$\begin{aligned} E(Y) &= \mu_y = 0 \times \Pr(Y = 0) + 1 \times \Pr(Y = 1) \\ &= 0 \times 0.050 + 1 \times 0.950 = 0.950. \end{aligned}$$

(b)

$$\begin{aligned} \text{Unemployment Rate} &= \frac{\#(\text{unemployed})}{\#(\text{labor force})} \\ &= \Pr(Y = 0) = 0.050 = 1 - 0.950 = 1 - E(Y). \end{aligned}$$

(c) Calculate the conditional probabilities first:

$$\Pr(Y = 0|X = 0) = \frac{\Pr(X = 0, Y = 0)}{\Pr(X = 0)} = \frac{0.045}{0.754} = 0.0597,$$

$$\Pr(Y = 1|X = 0) = \frac{\Pr(X = 0, Y = 1)}{\Pr(X = 0)} = \frac{0.709}{0.754} = 0.9403,$$

$$\Pr(Y = 0|X = 1) = \frac{\Pr(X = 1, Y = 0)}{\Pr(X = 1)} = \frac{0.005}{0.246} = 0.0203,$$

$$\Pr(Y = 1|X = 1) = \frac{\Pr(X = 1, Y = 1)}{\Pr(X = 1)} = \frac{0.241}{0.246} = 0.9797.$$

The conditional expectations are

$$\begin{aligned} E(Y|X = 1) &= 0 \times \Pr(Y = 0|X = 1) + 1 \times \Pr(Y = 1|X = 1) \\ &= 0 \times 0.0203 + 1 \times 0.9797 = 0.9797, \end{aligned}$$

$$\begin{aligned} E(Y|X = 0) &= 0 \times \Pr(Y = 0|X = 0) + 1 \times \Pr(Y = 1|X = 0) \\ &= 0 \times 0.0597 + 1 \times 0.9403 = 0.9403. \end{aligned}$$

- (d) Use the solution to part (b),

$$\begin{aligned} & \text{Unemployment rate for college grads} \\ & = 1 - E(Y|X=1) = 1 - 0.9797 = 0.0203. \end{aligned}$$

$$\begin{aligned} & \text{Unemployment rate for non-college grads} \\ & = 1 - E(Y|X=0) = 1 - 0.9403 = 0.0597. \end{aligned}$$

- (e) The probability that a randomly selected worker who is reported being unemployed is a college graduate is

$$\Pr(X=1|Y=0) = \frac{\Pr(X=1, Y=0)}{\Pr(Y=0)} = \frac{0.005}{0.050} = 0.1.$$

The probability that this worker is a non-college graduate is

$$\Pr(X=0|Y=0) = 1 - \Pr(X=1|Y=0) = 1 - 0.1 = 0.9.$$

- (f) Educational achievement and employment status are not independent because they do not satisfy that, for all values of x and y ,

$$\Pr(Y=y|X=x) = \Pr(Y=y).$$

For example,

$$\Pr(Y=0|X=0) = 0.0597 \neq \Pr(Y=0) = 0.050.$$

7. Using obvious notation, $C = M + F$; thus $\mu_C = \mu_M + \mu_F$ and $\sigma_C^2 = \sigma_M^2 + \sigma_F^2 + 2 \operatorname{cov}(M, F)$. This implies

(a) $\mu_C = 40 + 45 = \$85,000$ per year.

(b) $\operatorname{cor}(M, F) = \frac{\operatorname{Cov}(M, F)}{\sigma_M \sigma_F}$, so that $\operatorname{Cov}(M, F) = \sigma_M \sigma_F \operatorname{cor}(M, F)$. Thus

$\operatorname{Cov}(M, F) = 12 \times 18 \times 0.80 = 172.80$, where the units are squared thousands of dollars per year.

(c) $\sigma_C^2 = \sigma_M^2 + \sigma_F^2 + 2 \operatorname{cov}(M, F)$, so that $\sigma_C^2 = 12^2 + 18^2 + 2 \times 172.80 = 813.60$, and

$$\sigma_C = \sqrt{813.60} = 28.524 \text{ thousand dollars per year.}$$

- (d) First you need to look up the current Euro/dollar exchange rate in the Wall Street Journal, the Federal Reserve web page, or other financial data outlet. Suppose that this exchange rate is e (say $e = 0.80$ euros per dollar); each 1\$ is therefore worth e E. The mean is therefore $e\mu_C$ (in units of thousands of euros per year), and the standard deviation is $e\sigma_C$ (in units of thousands of euros per year). The correlation is unit-free, and is unchanged.

8. $\mu_Y = E(Y) = 1$, $\sigma_Y^2 = \operatorname{var}(Y) = 4$. With $Z = \frac{1}{2}(Y-1)$,

$$\begin{aligned} \mu_Z &= E\left(\frac{1}{2}(Y-1)\right) = \frac{1}{2}(\mu_Y - 1) = \frac{1}{2}(1-1) = 0, \\ \sigma_Z^2 &= \operatorname{var}\left(\frac{1}{2}(Y-1)\right) = \frac{1}{4}\sigma_Y^2 = \frac{1}{4} \times 4 = 1. \end{aligned}$$

9.

		Value of Y					Probability Distribution of X
		14	22	30	40	65	
Value of X	1	0.02	0.05	0.10	0.03	0.01	0.21
	5	0.17	0.15	0.05	0.02	0.01	0.40
	8	0.02	0.03	0.15	0.10	0.09	0.39
Probability distribution of Y		0.21	0.23	0.30	0.15	0.11	1.00

(a) The probability distribution is given in the table above.

$$E(Y) = 14 \times 0.21 + 22 \times 0.23 + 30 \times 0.30 + 40 \times 0.15 + 65 \times 0.11 = 30.15$$

$$E(Y^2) = 14^2 \times 0.21 + 22^2 \times 0.23 + 30^2 \times 0.30 + 40^2 \times 0.15 + 65^2 \times 0.11 = 1127.23$$

$$\text{Var}(Y) = E(Y^2) - [E(Y)]^2 = 218.21$$

$$\sigma_Y = 14.77$$

(b) Conditional Probability of $Y|X=8$ is given in the table below

Value of Y				
14	22	30	40	65
0.02/0.39	0.03/0.39	0.15/0.39	0.10/0.39	0.09/0.39

$$E(Y|X=8) = 14 \times (0.02/0.39) + 22 \times (0.03/0.39) + 30 \times (0.15/0.39)$$

$$+ 40 \times (0.10/0.39) + 65 \times (0.09/0.39) = 39.21$$

$$E(Y^2|X=8) = 14^2 \times (0.02/0.39) + 22^2 \times (0.03/0.39) + 30^2 \times (0.15/0.39)$$

$$+ 40^2 \times (0.10/0.39) + 65^2 \times (0.09/0.39) = 1778.7$$

$$\text{Var}(Y) = 1778.7 - 39.21^2 = 241.65$$

$$\sigma_{Y|X=8} = 15.54$$

$$(c) E(XY) = (1 \times 14 \times 0.02) + (1 \times 22 \times 0.05) + \dots + (8 \times 65 \times 0.09) = 171.7$$

$$\text{Cov}(X, Y) = E(XY) - E(X)E(Y) = 171.7 - 5.33 \times 30.15 = 11.0$$

$$\text{Corr}(X, Y) = \text{Cov}(X, Y) / (\sigma_X \sigma_Y) = 11.0 / (5.46 \times 14.77) = 0.136$$

10. Using the fact that if $Y \sim N(\mu_Y, \sigma_Y^2)$ then $\frac{Y-\mu_Y}{\sigma_Y} \sim N(0, 1)$ and Appendix Table 1, we have

(a)

$$\Pr(Y \leq 3) = \Pr\left(\frac{Y-1}{2} \leq \frac{3-1}{2}\right) = \Phi(1) = 0.8413.$$

(b)

$$\begin{aligned} \Pr(Y > 0) &= 1 - \Pr(Y \leq 0) \\ &= 1 - \Pr\left(\frac{Y-3}{3} \leq \frac{0-3}{3}\right) = 1 - \Phi(-1) = \Phi(1) = 0.8413. \end{aligned}$$

(c)

$$\begin{aligned}\Pr(40 \leq Y \leq 52) &= \Pr\left(\frac{40-50}{5} \leq \frac{Y-50}{5} \leq \frac{52-50}{5}\right) \\ &= \Phi(0.4) - \Phi(-2) = \Phi(0.4) - [1 - \Phi(2)] \\ &= 0.6554 - 1 + 0.9772 = 0.6326.\end{aligned}$$

(d)

$$\begin{aligned}\Pr(6 \leq Y \leq 8) &= \Pr\left(\frac{6-5}{\sqrt{2}} \leq \frac{Y-5}{\sqrt{2}} \leq \frac{8-5}{\sqrt{2}}\right) \\ &= \Phi(2.1213) - \Phi(0.7071) \\ &= 0.9831 - 0.7602 = 0.2229.\end{aligned}$$

11. (a) 0.90
 (b) 0.05
 (c) 0.05
 (d) When $Y \sim \chi^2_{10}$, then $Y/10 \sim F_{10,\infty}$.
 (e) $Y = Z^2$, where $Z \sim N(0,1)$, thus $\Pr(Y \leq 1) = \Pr(-1 \leq Z \leq 1) = 0.32$.
12. (a) 0.05
 (b) 0.950
 (c) 0.953
 (d) The t_{df} distribution and $N(0, 1)$ are approximately the same when df is large.
 (e) 0.10
 (f) 0.01
13. (a) $E(Y^2) = \text{Var}(Y) + \mu_Y^2 = 1 + 0 = 1$; $E(W^2) = \text{Var}(W) + \mu_W^2 = 100 + 0 = 100$.
 (b) Y and W are symmetric around 0, thus skewness is equal to 0; because their mean is zero, this means that the third moment is zero.
 (c) The kurtosis of the normal is 3, so $3 = \frac{E(Y-\mu_Y)^4}{\sigma_Y^4}$; solving yields $E(Y^4) = 3$; a similar calculation yields the results for W .
 (d) First, condition on $X = 0$, so that $S = W$:
- $$E(S|X = 0) = 0; E(S^2|X = 0) = 100, E(S^3|X = 0) = 0, E(S^4|X = 0) = 3 \times 100^2.$$
- Similarly,
- $$E(S|X = 1) = 0; E(S^2|X = 1) = 1, E(S^3|X = 1) = 0, E(S^4|X = 1) = 3.$$
- From the large of iterated expectations
- $$\begin{aligned}E(S) &= E(S|X = 0) \times \Pr(X = 0) + E(S|X = 1) \times \Pr(X = 1) = 0 \\ E(S^2) &= E(S^2|X = 0) \times \Pr(X = 0) + E(S^2|X = 1) \times \Pr(X = 1) = 100 \times 0.01 + 1 \times 0.99 = 1.99 \\ E(S^3) &= E(S^3|X = 0) \times \Pr(X = 0) + E(S^3|X = 1) \times \Pr(X = 1) = 0 \\ E(S^4) &= E(S^4|X = 0) \times \Pr(X = 0) + E(S^4|X = 1) \times \Pr(X = 1) = 3 \times 100^2 \times 0.01 + 3 \times 1 \times 0.99 = 302.97\end{aligned}$$

(e) $\mu_S = E(S) = 0$, thus $E(S - \mu_S)^3 = E(S^3) = 0$ from part d. Thus skewness = 0.

Similarly, $\sigma_S^2 = E(S - \mu_S)^2 = E(S^2) = 1.99$, and $E(S - \mu_S)^4 = E(S^4) = 302.97$.

Thus, kurtosis = $302.97/(1.99^2) = 76.5$

14. The central limit theorem suggests that when the sample size (n) is large, the distribution of the sample average (\bar{Y}) is approximately $N(\mu_Y, \sigma_Y^2)$ with $\sigma_Y^2 = \frac{\sigma_y^2}{n}$. Given $\mu_Y = 100$, $\sigma_y^2 = 43.0$,

(a) $n = 100$, $\sigma_{\bar{Y}}^2 = \frac{\sigma_y^2}{n} = \frac{43}{100} = 0.43$, and

$$\Pr(\bar{Y} \leq 101) = \Pr\left(\frac{\bar{Y} - 100}{\sqrt{0.43}} \leq \frac{101 - 100}{\sqrt{0.43}}\right) \approx \Phi(1.525) = 0.9364.$$

(b) $n = 165$, $\sigma_{\bar{Y}}^2 = \frac{\sigma_y^2}{n} = \frac{43}{165} = 0.2606$, and

$$\begin{aligned}\Pr(\bar{Y} > 98) &= 1 - \Pr(\bar{Y} \leq 98) = 1 - \Pr\left(\frac{\bar{Y} - 100}{\sqrt{0.2606}} \leq \frac{98 - 100}{\sqrt{0.2606}}\right) \\ &\approx 1 - \Phi(-3.9178) = \Phi(3.9178) = 1.000 \text{ (rounded to four decimal places).}\end{aligned}$$

(c) $n = 64$, $\sigma_{\bar{Y}}^2 = \frac{\sigma_y^2}{n} = \frac{43}{64} = 0.6719$, and

$$\begin{aligned}\Pr(101 \leq \bar{Y} \leq 103) &= \Pr\left(\frac{101 - 100}{\sqrt{0.6719}} \leq \frac{\bar{Y} - 100}{\sqrt{0.6719}} \leq \frac{103 - 100}{\sqrt{0.6719}}\right) \\ &\approx \Phi(3.6599) - \Phi(1.2200) = 0.9999 - 0.8888 = 0.1111.\end{aligned}$$

15. (a)

$$\begin{aligned}\Pr(9.6 \leq \bar{Y} \leq 10.4) &= \Pr\left(\frac{9.6 - 10}{\sqrt{4/n}} \leq \frac{\bar{Y} - 10}{\sqrt{4/n}} \leq \frac{10.4 - 10}{\sqrt{4/n}}\right) \\ &= \Pr\left(\frac{9.6 - 10}{\sqrt{4/n}} \leq Z \leq \frac{10.4 - 10}{\sqrt{4/n}}\right)\end{aligned}$$

where $Z \sim N(0, 1)$. Thus,

$$(i) \quad n = 20; \quad \Pr\left(\frac{9.6 - 10}{\sqrt{4/n}} \leq Z \leq \frac{10.4 - 10}{\sqrt{4/n}}\right) = \Pr(-0.89 \leq Z \leq 0.89) = 0.63$$

$$(ii) \quad n = 100; \quad \Pr\left(\frac{9.6 - 10}{\sqrt{4/n}} \leq Z \leq \frac{10.4 - 10}{\sqrt{4/n}}\right) = \Pr(-2.00 \leq Z \leq 2.00) = 0.954$$

$$(iii) \quad n = 1000; \quad \Pr\left(\frac{9.6 - 10}{\sqrt{4/n}} \leq Z \leq \frac{10.4 - 10}{\sqrt{4/n}}\right) = \Pr(-6.32 \leq Z \leq 6.32) = 1.000$$

(b)

$$\begin{aligned}\Pr(10 - c \leq \bar{Y} \leq 10 + c) &= \Pr\left(\frac{-c}{\sqrt{4/n}} \leq \frac{\bar{Y} - 10}{\sqrt{4/n}} \leq \frac{c}{\sqrt{4/n}}\right) \\ &= \Pr\left(\frac{-c}{\sqrt{4/n}} \leq Z \leq \frac{c}{\sqrt{4/n}}\right).\end{aligned}$$

As n get large $\frac{c}{\sqrt{4/n}}$ gets large, and the probability converges to 1.

- (c) This follows from (b) and the definition of convergence in probability given in Key Concept 2.6.
16. There are several ways to do this. Here is one way. Generate n draws of Y, Y_1, Y_2, \dots, Y_n . Let $X_i = 1$ if $Y_i < 3.6$, otherwise set $X_i = 0$. Notice that X_i is a Bernoulli random variables with $\mu_x = \Pr(X = 1) = \Pr(Y < 3.6)$. Compute \bar{X} . Because \bar{X} converges in probability to $\mu_x = \Pr(X = 1) = \Pr(Y < 3.6)$, \bar{X} will be an accurate approximation if n is large.
17. $\mu_y = 0.4$ and $\sigma_y^2 = 0.4 \times 0.6 = 0.24$
- (a) (i) $P(\bar{Y} \geq 0.43) = \Pr\left(\frac{\bar{Y} - 0.4}{\sqrt{0.24/n}} \geq \frac{0.43 - 0.4}{\sqrt{0.24/n}}\right) = \Pr\left(\frac{\bar{Y} - 0.4}{\sqrt{0.24/n}} \geq 0.6124\right) = 0.27$
- (ii) $P(\bar{Y} \leq 0.37) = \Pr\left(\frac{\bar{Y} - 0.4}{\sqrt{0.24/n}} \leq \frac{0.37 - 0.4}{\sqrt{0.24/n}}\right) = \Pr\left(\frac{\bar{Y} - 0.4}{\sqrt{0.24/n}} \leq -1.22\right) = 0.11$
- (b) We know $\Pr(-1.96 \leq Z \leq 1.96) = 0.95$, thus we want n to satisfy $0.41 = \frac{0.41 - 0.4}{\sqrt{0.24/n}} > -1.96$ and $\frac{0.39 - 0.4}{\sqrt{0.24/n}} < -1.96$. Solving these inequalities yields $n \geq 9220$.
18. $\Pr(Y = \$0) = 0.95$, $\Pr(Y = \$20000) = 0.05$.

(a) The mean of Y is

$$\mu_y = 0 \times \Pr(Y = \$0) + 20,000 \times \Pr(Y = \$20000) = \$1000.$$

The variance of Y is

$$\begin{aligned}\sigma_y^2 &= E[(Y - \mu_y)^2] \\ &= (0 - 1000)^2 \times \Pr(Y = 0) + (20000 - 1000)^2 \times \Pr(Y = 20000) \\ &= (-1000)^2 \times 0.95 + 19000^2 \times 0.05 = 1.9 \times 10^7,\end{aligned}$$

so the standard deviation of Y is $\sigma_y = (1.9 \times 10^7)^{\frac{1}{2}} = \4359 .

(b) (i) $E(\bar{Y}) = \mu_y = \$1000$, $\sigma_{\bar{Y}}^2 = \frac{\sigma_y^2}{n} = \frac{1.9 \times 10^7}{100} = 1.9 \times 10^5$.

(ii) Using the central limit theorem,

$$\begin{aligned}\Pr(\bar{Y} > 2000) &= 1 - \Pr(\bar{Y} \leq 2000) \\ &= 1 - \Pr\left(\frac{\bar{Y} - 1000}{\sqrt{1.9 \times 10^5}} \leq \frac{2,000 - 1,000}{\sqrt{1.9 \times 10^5}}\right) \\ &\approx 1 - \Phi(2.2942) = 1 - 0.9891 = 0.0109.\end{aligned}$$

19. (a)

$$\begin{aligned}\Pr(Y = y_j) &= \sum_{i=1}^l \Pr(X = x_i, Y = y_j) \\ &= \sum_{i=1}^l \Pr(Y = y_j | X = x_i) \Pr(X = x_i)\end{aligned}$$

(b)

$$\begin{aligned}E(Y) &= \sum_{j=1}^k y_j \Pr(Y = y_j) = \sum_{j=1}^k y_j \sum_{i=1}^l \Pr(Y = y_j | X = x_i) \Pr(X = x_i) \\ &= \sum_{i=1}^l \left(\sum_{j=1}^k y_j \Pr(Y = y_j | X = x_i) \right) \Pr(X = x_i) \\ &= \sum_{i=1}^l E(Y | X = x_i) \Pr(X = x_i).\end{aligned}$$

(c) When X and Y are independent,

$$\Pr(X = x_i, Y = y_j) = \Pr(X = x_i) \Pr(Y = y_j),$$

so

$$\begin{aligned}\sigma_{XY} &= E[(X - \mu_X)(Y - \mu_Y)] \\ &= \sum_{i=1}^l \sum_{j=1}^k (x_i - \mu_X)(y_j - \mu_Y) \Pr(X = x_i, Y = y_j) \\ &= \sum_{i=1}^l \sum_{j=1}^k (x_i - \mu_X)(y_j - \mu_Y) \Pr(X = x_i) \Pr(Y = y_j) \\ &= \left(\sum_{i=1}^l (x_i - \mu_X) \Pr(X = x_i) \right) \left(\sum_{j=1}^k (y_j - \mu_Y) \Pr(Y = y_j) \right) \\ &= E(X - \mu_X) E(Y - \mu_Y) = 0 \times 0 = 0,\end{aligned}$$

$$cor(X, Y) = \frac{\sigma_{XY}}{\sigma_X \sigma_Y} = \frac{0}{\sigma_X \sigma_Y} = 0.$$

20. (a) $\Pr(Y = y_i) = \sum_{j=1}^l \sum_{h=1}^m \Pr(Y = y_i | X = x_j, Z = z_h) \Pr(X = x_j, Z = z_h)$

(b)

$$\begin{aligned}E(Y) &= \sum_{i=1}^k y_i \Pr(Y = y_i) \Pr(Y = y_i) \\ &= \sum_{i=1}^k y_i \sum_{j=1}^l \sum_{h=1}^m \Pr(Y = y_i | X = x_j, Z = z_h) \Pr(X = x_j, Z = z_h) \\ &= \sum_{j=1}^l \sum_{h=1}^m \left[\sum_{i=1}^k y_i \Pr(Y = y_i | X = x_j, Z = z_h) \right] \Pr(X = x_j, Z = z_h) \\ &= \sum_{j=1}^l \sum_{h=1}^m E(Y | X = x_j, Z = z_h) \Pr(X = x_j, Z = z_h)\end{aligned}$$

where the first line in the definition of the mean, the second uses (a), the third is a rearrangement, and the final line uses the definition of the conditional expectation.

21. (a)

$$\begin{aligned} E(X - \mu)^3 &= E[(X - \mu)^2(X - \mu)] = E[X^3 - 2X^2\mu + X\mu^2 - X^2\mu + 2X\mu^2 - \mu^3] \\ &= E(X^3) - 3E(X^2)\mu + 3E(X)\mu^2 - \mu^3 = E(X^3) - 3E(X^2)E(X) + 3E(X)[E(X)]^2 - [E(X)]^3 \\ &= E(X^3) - 3E(X^2)E(X) + 2E(X)^3 \end{aligned}$$

(b)

$$\begin{aligned} E(X - \mu)^4 &= E[(X^3 - 3X^2\mu + 3X\mu^2 - \mu^3)(X - \mu)] \\ &= E[X^4 - 3X^3\mu + 3X^2\mu^2 - X\mu^3 - X^3\mu + 3X^2\mu^2 - 3X\mu^3 + \mu^4] \\ &= E(X^4) - 4E(X^3)E(X) + 6E(X^2)E(X)^2 - 4E(X)E(X)^3 + E(X)^4 \\ &= E(X^4) - 4[E(X)][E(X^3)] + 6[E(X)]^2[E(X^2)] - 3[E(X)]^4 \end{aligned}$$

22. The mean and variance of R are given by

$$\begin{aligned} \mu &= w \times 0.08 + (1-w) \times 0.05 \\ \sigma^2 &= w^2 \times 0.07^2 + (1-w)^2 \times 0.042 + 2 \times w \times (1-w) \times [0.07 \times 0.04 \times 0.25] \end{aligned}$$

where $0.07 \times 0.04 \times 0.25 = \text{Cov}(R_s, R_b)$ follows from the definition of the correlation between R_s and R_b .

- (a) $\mu = 0.065$; $\sigma = 0.044$
- (b) $\mu = 0.0725$; $\sigma = 0.056$
- (c) $w = 1$ maximizes μ ; $\sigma = 0.07$ for this value of w .
- (d) The derivative of σ^2 with respect to w is

$$\begin{aligned} \frac{d\sigma^2}{dw} &= 2w \times 0.07^2 - 2(1-w) \times 0.04^2 + (2-4w) \times [0.07 \times 0.04 \times 0.25] \\ &= 0.0102w - 0.0018 \end{aligned}$$

solving for w yields $w = 18/102 = 0.18$. (Notice that the second derivative is positive, so that this is the global minimum.) With $w = 0.18$, $\sigma_R = .038$.

23. X and Z are two independently distributed standard normal random variables, so

$$\mu_X = \mu_Z = 0, \sigma_X^2 = \sigma_Z^2 = 1, \sigma_{XZ} = 0.$$

- (a) Because of the independence between X and Z , $\Pr(Z = z | X = x) = \Pr(Z = z)$, and $E(Z|X) = E(Z) = 0$. Thus $E(Y|X) = E(X^2 + Z|X) = E(X^2|X) + E(Z|X) = X^2 + 0 = X^2$.
- (b) $E(X^2) = \sigma_X^2 + \mu_X^2 = 1$, and $\mu_Y = E(X^2 + Z) = E(X^2) + \mu_Z = 1 + 0 = 1$.
- (c) $E(XY) = E(X^3 + ZX) = E(X^3) + E(ZX)$. Using the fact that the odd moments of a standard normal random variable are all zero, we have $E(X^3) = 0$. Using the independence between X and Z , we have $E(ZX) = \mu_Z \mu_X = 0$. Thus $E(XY) = E(X^3) + E(ZX) = 0$.

(d)

$$\begin{aligned} \text{Cov}(XY) &= E[(X - \mu_X)(Y - \mu_Y)] = E[(X - 0)(Y - 1)] \\ &= E(XY - X) = E(XY) - E(X) \\ &= 0 - 0 = 0. \end{aligned}$$

$$\text{cor}(X, Y) = \frac{\sigma_{XY}}{\sigma_X \sigma_Y} = \frac{0}{\sigma_X \sigma_Y} = 0.$$

24. (a) $E(Y_i^2) = \sigma^2 + \mu^2 = \sigma^2$ and the result follows directly.

(b) (Y/σ) is distributed i.i.d. $N(0,1)$, $W = \sum_{i=1}^n (Y_i/\sigma)^2$, and the result follows from the definition of a χ_n^2 random variable.

$$(c) E(W) = E(W) = E \sum_{i=1}^n \frac{Y_i^2}{\sigma^2} = \sum_{i=1}^n E \frac{Y_i^2}{\sigma^2} = n.$$

(d) Write

$$V = \frac{Y_1}{\sqrt{\frac{\sum_{i=2}^n Y_i^2}{n-1}}} = \frac{Y_1/\sigma}{\sqrt{\frac{\sum_{i=2}^n (Y_i/\sigma)^2}{n-1}}}$$

which follows from dividing the numerator and denominator by σ . $Y_1/\sigma \sim N(0,1)$, $\sum_{i=2}^n (Y_i/\sigma)^2 \sim \chi_{n-1}^2$, and Y_1/σ and $\sum_{i=2}^n (Y_i/\sigma)^2$ are independent. The result then follows from the definition of the t distribution.